# Error-Rate Prediction for Probabilistic Circuits with More General Structures 

Mark S. K. Lau, Keck-Voon Ling, Arun Bhanu, and Vincent J. Mooney III*<br>School of Electrical and Electronic Engineering<br>Nanyang Technological University<br>50 Nanyang Avenue, Singapore 639798<br>\{marklausk, ekvling, arunbhanu, vjmooney\}@ntu.edu.sg


#### Abstract

A methodology has been proposed recently to predict error-rates of probabilistic circuits having a cascade structure. It was able to predict reasonably accurately for probabilistic ripple-carry and carry-skip adders. The objective of the present paper is twofold. First, the methodology is applied, for the first time in the literature, to a probabilistic carryselect adder, which has a more complex structure than the adders mentioned above. This is to provide additional evidence that the method is versatile and applicable to some non-trivial circuits. Second, the present paper shows that the methodology is also applicable to some seemingly non-cascade circuits. The key technique is to appropriately group circuit components into various blocks before applying the methodology. Such a preprocessing may potentially widen the scope of applicability of the methodology.


## I. Introduction

As the feature size of CMOS continues to shrink, CMOS circuits become increasingly susceptible to various kinds of errors, failures, and process-induced variations. Instead of spending disproportionate efforts to avoid errors completely, probabilistic computing proposes to embrace occasional errors caused by erroneous chips [1]. Manifested in probabilistic computing are potential trade-offs between correctness of circuit operation and design parameters. A good trade-off can possibly result in significant power savings $[2,3]$ or increased circuit speeds [4].

For systematic design, performance evaluation, and resource allocation of probabilistic circuits, accurate prediction of error-rates is essential. A method was proposed in [5] to predict error-rates of probabilistic circuits that are cascades of some circuit blocks as shown in Fig. 1. Such a cascade structure can be found in, for examples, various adder circuits such as ripple-carry and carry-skip (CSAs) adders. In [5] the method was used to predict error-rates of a probabilistic CSA. The predicted error-rates were reasonably close to that obtained from HSPICE simulations.

A contribution of the present paper is to apply the method of [5] to a more complex circuit, namely, a proba-

[^0]

Fig. 1. A cascade circuit to which the method of [5] is applicable.


Fig. 2. Transforming converging fan-outs into a cascade structure.
bilistic carry-select adder (PCSLA). The PCSLA has two multi-bit ripple-carry adders working in parallel, and has a more complex circuit structure (e.g., converging fanouts) than the CSA studied in [5]. The results shown in the present paper can provide additional evidence that the method is versatile and applicable to non-trivial circuits.

In a recent study, we observed that some circuits of a directed-acyclic-graph structure can be re-configured into the same cascade structure as Fig. 1. It was done by properly grouping circuit components into blocks. We have seen in our study that this technique works for a Wallace tree multiplier. Motivated by the study, as the second contribution of the present paper, we suggest that the method of [5] is also applicable to some circuits with seemingly non-cascade structures; the key is a preprocessing step that can reconfigure components of a target circuit into blocks properly.

An example is shown in Fig. 2 (a), where the error-rate of $z$ is to be determined. At first, it seems that this circuit does not have a cascade structure because there are two fan-out signals converging at E . One may be tempted to
decompose the circuit into three cascade sub-circuits, as shown in Fig. 2 (b). The method of [5] is then applied to sub-circuits 1 and 2 separately to obtain probabilities of $x$ and $y$, denoted as $p_{x}$ and $p_{y}$. If $x$ and $y$ are statistically independent, a joint probability of $x$ and $y$ is simply the product $p_{x} p_{y}$. Having this joint probability, one can apply the method of [5] again to sub-circuit 3 to obtain the errorrate of $z$.

However, this technique may fail when $x$ and $y$ are correlated. Unless the correlation is known in advance, there may be insufficient information to estimate a joint probability of $x$ and $y$ since $p_{x}$ and $p_{y}$ were estimated separately.

This is now tackled by appropriately re-grouping the components into blocks, i.e., preprocessing, before applying the method. We group A and B as a block, C and D as a block, while E itself is another block, as shown in Fig. 2 (c). It is now clear that Fig. 2 (c) exhibits the same cascade structure as Fig. 1, and thus the method of [5] can be applied. Joint probabilities of $x$ and $y$ are now given directly by the method since they are outputs from the same block. Similar ideas were used for another error-rate predicting method in [6].

In view of this example, we are able to apply the method of [5] to circuits having cascade as well as some seemingly non-cascade structures. This technique is demonstrated in the present paper with a PCSLA as an example.

## II. Carry-select Adders

We consider a carry-select adder (CSLA) which is a cascade of several identical multi-bit adders, as depicted in Fig. 3 (a). The internal structure of the $i$ th multi-bit adder is shown in Fig. 4. For a detailed discussion of CSLAs, see for example pg. 192 of [7] and pg. 575 of [8].

The multi-bit adder shown in Fig. 4 accepts a carryin $c_{i, 1}$ and two segments of addend bits. The addend segments are added, and the resulting carry-out $c_{i+1,1}$ will serve as the carry-in of the next multi-bit adder.

Before a carry-in $c_{i, 1}$ arrives, the multi-bit adder performs major amount of work required by the addition. The multi-bit adder performs the addition twice with two possible carry-in values (VSS and VDD); the additions are performed in parallel by two ripple-carry adders. Once an actual carry-in has arrived, the multiplexers (MUXes) will select the correct addition result based on the value of the actual carry-in. Then, the correct addition result will be ready for output almost instantly.

We decompose a CSLA in two different ways. The first way, which is perhaps more obvious, is to view a CSLA as a cascade of multi-bit adders, as shown in Fig. 3 (a). The CSLA then exhibits the same cascade structure of Fig. 1. We are allowed to apply the method of [5] directly, considering each multi-bit adder as a single block. This is a straightforward way of decomposing a CSLA. However, the method becomes less efficient for a block of a long bit-length [5], as the complexity of characterizing a block grows rapidly with its number of inputs/outputs.


Fig. 3. Two representations of the same CSLA.


Fig. 4. A CSLA is a cascade of instances of this multi-bit adder.


Fig. 5. The multi-bit adder of Fig. 4 is a cascade of three types of primitive blocks.

Such a scalability issue can be alleviated by further dividing a multi-bit adder into a cascade of sub-circuits, where each sub-circuit has only a small number of inputs/outputs. The multi-bit adder of Fig. 4, however, has no apparent cascade structure if we assign each fulladder (FA) or multiplexer (MUX) an individual block. It is then not clear how the multi-bit adder of Fig. 4 can be sub-divided into a cascade of smaller blocks.

In this paper, we observe that a cascade structure would immediately emerge if we group the FAs and MUXes into three types of blocks, as shown in Fig. 5. This results in another cascade representation of a CSLA, as depicted in Fig. 3 (b). This cascade structure can be handled by the method of [5]. The above-mentioned scalability issue does
not exist any more. This is because each block of Fig. 5 corresponds to no more than a single bit of a CSLA, and hence the number of inputs/outputs of a block does not grow with the bit-length of a multi-bit adder.

The above discussions illustrate that, by appropriately grouping circuit components into blocks, circuits of seemingly non-cascade structure (e.g., the multi-bit adder of Fig. 4) can also be partitioned to form a cascade structure that can be handled by the method of [5].

## III. A Framework for Error-rate Prediction

In this section, we briefly present the method of [5] and the associated equations for error-rate prediction. Discussions here are self-contained, and can be understood without referring to the original paper [5].

The method deals with a cascade of $N$ circuit blocks shown in Fig. 1. Block $i$ has $m_{i}+q_{i}$ inputs and $n_{i}+q_{i+1}$ outputs; the inputs $\left(c_{i, 1}, \ldots, c_{i, q_{i}}\right)$ and $\left(a_{i, 1}, \ldots, a_{i, m_{i}}\right)$ are statistically independent. Internal structures of the circuit blocks need not to be specified in advance, and there are no strong assumptions imposed.

## A. Notation and Mathematical Model

We use the notation shown in Fig. 1. The first subscript refers to the block number and the second subscript refers to the bit position within a block. We shall sometimes use vector notation, which is denoted by bold symbols such as $\boldsymbol{a}_{i}$ and $\boldsymbol{c}_{i}$. The definitions of some symbols are given in Table I for the ease of reference. With vector notation, we present the $i$ th block in a compact manner in Fig. 6 (a).

We introduce two functions, $f_{i j}^{c}$ and $f_{i j}^{s}$, that model the input-output relations of the block in Fig. 6 (a). Namely,

$$
c_{i+1, j}=f_{i j}^{c}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right), \quad s_{i j}=f_{i j}^{s}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right) .
$$

We also write in vector forms:

$$
\begin{aligned}
\boldsymbol{c}_{i+1} & =\left(c_{i+1,1}, c_{i+1,2}, \ldots, c_{i+1, q_{i+1}}\right) \\
& =\boldsymbol{f}_{i}^{c}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right) \stackrel{\text { def }}{=}\left(f_{i, 1}^{c}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right), \ldots, f_{i, q_{i+1}}^{c}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right)\right) \\
\boldsymbol{s}_{i} & =\left(s_{i, 1}, s_{i, 2}, \ldots, s_{i, n_{i}}\right) \\
& =\boldsymbol{f}_{i}^{s}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right) \stackrel{\text { def }}{=}\left(f_{i, 1}^{s}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right), \ldots, f_{i, n_{i}}^{s}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right)\right)
\end{aligned}
$$

Figure 6 (b) shows a probabilistic block with inputs and outputs $\boldsymbol{c}_{i}^{\prime}, \boldsymbol{c}_{i+1}^{\prime}, \boldsymbol{a}_{i}$ and $\boldsymbol{s}_{i}^{\prime}$. We assume that Figs. 6 (a) and (b) have the same input $\boldsymbol{a}_{i}$. Hence, the probabilistic block's input $\boldsymbol{a}$ is always correct. However, $\boldsymbol{c}_{i}$ and $\boldsymbol{c}_{i}^{\prime}$ may be different, and we use a prime to indicate the distinction. The prime notation here does not denote complementation, and hence variables such as $\boldsymbol{c}_{i}^{\prime}$ and $\boldsymbol{c}_{i}$ are in general completely distinct.

Assume that errors at the outputs of a probabilistic block are caused by some inherent error sources. The error sources cause an error and subsequently flip an output bit. Variables $e_{i j}^{c}$ and $e_{i j}^{s}$ were introduced in [5] such that

$$
c_{i+1, j}^{\prime}=e_{i j}^{c} \oplus f_{i j}^{c}\left(\boldsymbol{c}_{i}^{\prime}, \boldsymbol{a}_{i}\right), \quad s_{i, j}^{\prime}=e_{i j}^{s} \oplus f_{i j}^{s}\left(\boldsymbol{c}_{i}^{\prime}, \boldsymbol{a}_{i}\right)
$$

TABLE I
Mathematical symbols for the $i$ Th block.

| Types | Symbols for the $i$ th block |
| ---: | :--- |
| Input | $\boldsymbol{a}_{i}=\left(a_{i, 1}, a_{i, 2}, \ldots, a_{i, m_{i}}\right)$ |
| Input | $\boldsymbol{c}_{i}=\left(c_{i, 1}, c_{i, 2}, \ldots, c_{i, q_{i}}\right)$ |
| Output | $\boldsymbol{c}_{i+1}=\left(c_{i+1,1}, \ldots, c_{i+1, q_{i+1}}\right)$ |
| Probabilistic input | $\boldsymbol{c}_{i}^{\prime}=\left(c_{i, 1}^{\prime}, c_{i, 2}^{\prime}, \ldots, c_{i, q_{i}}^{\prime}\right)$ |
| Probabilistic output | $\boldsymbol{c}_{i+1}^{\prime}=\left(c_{i+1,1}^{\prime}, \ldots, c_{i+1, q_{i+1}}^{\prime}\right)$ |
| Output bit-flip indicator | $\boldsymbol{e}_{i}^{c}=\left(e_{i, 1}^{c}, e_{i, 2}^{c}, \ldots, e_{i, q_{i+1}}^{c}\right)$ |
| Output | $\boldsymbol{s}_{i}=\left(s_{i, 1}, s_{i, 2}, \ldots, s_{i, n_{i}}\right)$ |
| Probabilistic output | $\boldsymbol{s}_{i}^{\prime}=\left(s_{i, 1}^{\prime}, s_{i, 2}^{\prime}, \ldots, s_{i, n_{i}}^{\prime}\right)$ |
| Output bit-flip indicator | $\boldsymbol{e}_{i}^{s}=\left(e_{i, 1}^{s}, e_{i, 2}^{s}, \ldots, e_{i, n_{i}}^{s}\right)$ |



Fig. 6. Characterizing a probabilistic block.

Hence, for example, $e_{i j}^{c}=1$ means that the $i$ th probabilistic block has made an error at the $j$ th carry output. The above equations can also be written in vector forms:

$$
\begin{equation*}
\boldsymbol{c}_{i+1}^{\prime}=\boldsymbol{e}_{i}^{c} \oplus \boldsymbol{f}_{i}^{c}\left(\boldsymbol{c}_{i}^{\prime}, a_{i}\right), \quad \boldsymbol{s}_{i}^{\prime}=\boldsymbol{e}_{i}^{s} \oplus \boldsymbol{f}_{i}^{s}\left(\boldsymbol{c}_{i}^{\prime}, \boldsymbol{a}_{i}\right) \tag{1}
\end{equation*}
$$

where $\boldsymbol{e}_{i}^{c}$ and $\boldsymbol{e}_{i}^{s}$ were defined in Table I, and $\oplus$ means component-wise XOR on two vectors of binary numbers.

Finally, two assumptions are required, as stated in [5]. First, $\boldsymbol{c}_{i}$ and $\boldsymbol{a}_{i}$ are statistically independent. Second, $\boldsymbol{e}_{i}^{c}$ is conditionally independent of $\boldsymbol{c}_{i}$ given $\boldsymbol{c}_{i+1}^{\prime}$ and $\boldsymbol{a}_{i}$.

## B. Equations for Error-rate Prediction

Suppose that one is interested in predicting the probability that $\boldsymbol{c}_{i+1}=\boldsymbol{\alpha}, \boldsymbol{c}_{i+1}^{\prime}=\boldsymbol{\beta}$ for some vectors $\boldsymbol{\alpha}$ and $\boldsymbol{\beta}$ of binary numbers. Let $\boldsymbol{\gamma}^{c}, \boldsymbol{\delta}, \boldsymbol{\epsilon}$ and $\boldsymbol{\zeta}$ be vectors of binary numbers, and let $\mathcal{C}(\boldsymbol{\alpha}, \boldsymbol{\beta})$ be the set of all $\left(\boldsymbol{\gamma}^{c}, \boldsymbol{\delta}, \boldsymbol{\epsilon}, \boldsymbol{\zeta}\right)$ such that $\boldsymbol{e}_{i}^{c}=\boldsymbol{\gamma}^{c}, \boldsymbol{c}_{i}=\boldsymbol{\delta}, \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}$ and $\boldsymbol{a}_{i}=\boldsymbol{\zeta}$ results in

$$
\begin{aligned}
c_{i+1} & =\boldsymbol{f}_{i}^{c}\left(\boldsymbol{c}_{i}, \boldsymbol{a}_{i}\right)=\boldsymbol{f}_{i}^{c}(\boldsymbol{\delta}, \boldsymbol{\zeta})=\boldsymbol{\alpha} \\
\boldsymbol{c}_{i+1}^{\prime} & =\boldsymbol{e}_{i}^{c} \oplus \boldsymbol{f}_{i}^{c}\left(\boldsymbol{c}_{i}^{\prime}, \boldsymbol{a}_{i}\right)=\boldsymbol{\gamma}^{c} \oplus \boldsymbol{f}_{i}^{c}(\boldsymbol{\epsilon}, \boldsymbol{\zeta})=\boldsymbol{\beta}
\end{aligned}
$$

The set $\mathcal{C}(\boldsymbol{\alpha}, \boldsymbol{\beta})$ can be obtained by, for example, exhausting all possible values of $\boldsymbol{e}_{i}^{c}, \boldsymbol{c}_{i}, \boldsymbol{c}_{i}^{\prime}$ and $\boldsymbol{a}_{i}$.

Equation (2) in the next page was derived in [5]. It is recursive, namely, a probability for the $(i+1)$ th block, i.e., $P\left(\boldsymbol{c}_{i+1}^{\prime}=\boldsymbol{\alpha}, \boldsymbol{c}_{i+1}=\boldsymbol{\beta}\right)$, is computed from probabilities

$$
\begin{align*}
P\left(\boldsymbol{c}_{i+1}^{\prime}=\boldsymbol{\alpha}, \boldsymbol{c}_{i+1}=\boldsymbol{\beta}\right) & =\sum_{\mathcal{C}(\boldsymbol{\alpha}, \boldsymbol{\beta})} P\left(\boldsymbol{e}_{i}^{c}=\gamma^{c} \mid \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) P\left(\boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) P\left(\boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{c}_{i}=\boldsymbol{\delta}\right) .  \tag{2}\\
P\left(s_{i j}^{\prime} \neq s_{i j}\right) & =\sum_{\mathcal{S}_{i j}} P\left(e_{i j}^{s}=\gamma^{s} \mid \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) P\left(\boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) P\left(\boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{c}_{i}=\boldsymbol{\delta}\right) \tag{3}
\end{align*}
$$

of the previous block denoted as $P\left(\boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{c}_{i}=\boldsymbol{\delta}\right)$. If probabilities of the first block are known, (2) can be used iteratively to obtain probabilities of the other blocks. In (2), $P\left(\boldsymbol{a}_{i}=\boldsymbol{\zeta}\right)$ is a probability of the data, which is given by the data statistics and is known in advance.

The conditional probability in (2) indeed characterizes the probabilistic behavior of the $i$ th block. Recall from (1) that $\boldsymbol{e}_{i}^{c}$ indicates whether errors were made at the output $\boldsymbol{c}_{i}$ of the $i$ block. Therefore, one may think of the conditional probability as a probability of error given the inputs $\boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}$ and $\boldsymbol{a}_{i}=\boldsymbol{\zeta}$. It characterizes the response of the $i$ th probabilistic block to a particular input pattern.

The conditional probability in (2) can be estimated by experiments on the $i$ th probabilistic block, where a large number of randomly generated inputs are fed to the block and the number of erroneous outputs are recorded subsequently [5]. This will be discussed in the next subsection.

Similarly, error-rates for $s_{i}$ can be predicted by equation (3), whose derivation is found in [5]. Suppose that one wishes to estimate the probability that $s_{i j} \neq s_{i j}^{\prime}$. Let $\gamma^{\boldsymbol{s}}$ be a binary number, and let $\boldsymbol{\delta}, \boldsymbol{\epsilon}$ and $\boldsymbol{\zeta}$ be vectors of binary numbers. Let $\mathcal{S}_{i j}$ be the set of all possible ( $\gamma^{s}, \boldsymbol{\delta}, \boldsymbol{\epsilon}, \boldsymbol{\zeta}$ ) such that $s_{i j}^{\prime} \neq s_{i j}$ if $e_{i j}^{s}=\gamma^{s}, \boldsymbol{c}_{i}=\boldsymbol{\delta}, \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}$ and $\boldsymbol{a}_{i}=\boldsymbol{\zeta}$.

## C. Characterizing Probabilistic Blocks

Recall that (2) and (3) require estimates of the following probabilities for the $i$ th block:

$$
\begin{equation*}
P\left(\boldsymbol{e}_{i}^{c}=\gamma^{c} \mid \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{a}_{i}=\boldsymbol{\zeta}\right), P\left(e_{i j}^{s}=\gamma^{s} \mid \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) \tag{4}
\end{equation*}
$$

It is done using the experiment illustrated in Fig. 6 (c). The lower block is the $i$ th block of a probabilistic circuit, and the upper one is deterministic. This process is referred to as characterizing a probabilistic block [5, 9, 10].

The two blocks in Fig. 6 (c) have the same inputs generated by random-number generators of certain probability distributions. Independence between $\boldsymbol{c}_{\mathrm{in}}$ and $\boldsymbol{a}$ is needed, due to the first assumption mentioned in the last paragraph of Section III-A. Outputs given by the blocks are compared using the XOR function to check for errors.

To estimate (4), we perform the experiment of Fig. 6 (c) with $\Theta$ sets of randomly generated inputs, where each set contains a vector of $q_{i}$ bits for $\boldsymbol{c}_{\mathrm{in}}$, and a vector of $m_{i}$ bits for $\boldsymbol{a}$. Let $\Gamma(\Theta)$ be the number of randomly generated input sets such that $\boldsymbol{c}_{\mathrm{in}}=\boldsymbol{\epsilon}$ and $\boldsymbol{a}=\boldsymbol{\zeta}$.

Out of those $\Gamma(\Theta)$ input sets, there are $\Lambda^{c}(\Theta)$ sets causing the output of the upper XOR function equal to $\gamma^{c}$; and there are $\Lambda^{s}(\Theta)$ sets causing the $j$ th bit of the output of the lower XOR function equal to $\gamma^{s}$. By the strong law


Fig. 7. A probabilistic block used in HSPICE simulations.
of large numbers (Theorem 4.1 of [11] and equation (4.1) thereafter), we do the following estimation:

$$
\begin{align*}
& P\left(\boldsymbol{e}_{i}^{c}=\boldsymbol{\gamma}^{c} \mid \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) \approx \lim _{\Theta \rightarrow \infty} \frac{\Lambda^{c}(\Theta)}{\Gamma(\Theta)}  \tag{5}\\
& P\left(e_{i j}^{s}=\gamma^{s} \mid \boldsymbol{c}_{i}^{\prime}=\boldsymbol{\epsilon}, \boldsymbol{a}_{i}=\boldsymbol{\zeta}\right) \approx \lim _{\Theta \rightarrow \infty} \frac{\Lambda^{s}(\Theta)}{\Gamma(\Theta)} \tag{6}
\end{align*}
$$

In practice, an infinitely large $\Theta$ is impossible. However, one has to make sure that $\Theta$ is sufficiently large to allow the estimated probabilities to converge and stabilize.

## IV. Application to Probabilistic CSLA

The method of [5] described previously is now applied to a PCSLA. In particular, we use (2) and (3) to predict error-rates of a PCSLA, and then compare with errorrates obtained from HSPICE simulations. We also experiment with two decomposition methods proposed in Section II, and discuss the accuracy of their predictions.

## A. Obtaining Error-rates by HSPICE Simulation

Following $[2,5,10,9,12]$, we construct a probabilistic circuit by coupling noise sources to the output terminals of some deterministic components. This is seen from Fig. 7, which shows a diagram of a block in our simulations.

Each of the noise sources is generated independently, and is of Gaussian distribution with zero mean and nonzero root-mean-square (RMS) value. A noise source is simulated in HSPICE by a voltage-controlled voltagesource (VCVS) with a voltage gain equal to the desired noise RMS. In Fig. 7 VCVSes are labeled by E1, E2, ..., etc. Random numbers of the standard Gaussian distribution are generated by Matlab. A VCVS multiplies these random numbers by its voltage gain to get a Gaussian
noise source of the desired RMS. Noise samples are added to the output terminals every 500 ps .

Our simulations use the Synopsys 90 nm generic library. For this technology, the nominal voltage is 1.2 V . We follow $[5,9,10,12]$ to use a noise RMS of 0.2 V .

To construct a PCSLA of 16 bits, we cascade four identical copies of the block shown in Fig. 7. Two hundred thousand sets of input bits are randomly generated by Matlab. All input bits are independent of each other, and are of uniform distribution. The sampling period is 5 ns , which is longer than the worst-case propagation delays of the LSB-input-to-MSB-sum critical paths of the 16-bit PCSLA. That is, every 5 ns , a newly generated set of input bits is fed to the 16 -bit PCSLA, and at the same time, the output bits (which are due to the past input vectors) are sampled. All blocks in the 16 -bit PCSLA are identical, and have the same supply voltages chosen from the set $\{0.8,0.9, \ldots, 1.2\}$.

A sampled output has the logical state zero (one) if the output voltage is lower (equal to or higher) than half of the supply voltage. We compare the sampled outputs with the correct results. In our simulations, instead of actually implementing a deterministic CSLA in HSPICE, the deterministic (correct) outputs are calculated by arithmetic functions. The simulated error-rate of the $i$ th sum-bit of a PCSLA is defined as the number of errors observed at the $i$ th sum-bit in the HSPICE simulation, normalized by the number of observed output samples, i.e., 200,000.

## B. Predicting Error-rates

In the following, we are going to apply (2) and (3) to predict error-rates of a PCSLA, which are then compared with those error-rates obtained by HSPICE simulations.

As suggested in Section II, we may consider a PCSLA as a cascade of identical instances of the block in Fig. 4. In this case, the block has to be characterized by carrying out an experiment illustrated in Fig. 6 (c). Since all blocks are identical in the PCSLA considered here, the characterization has to be done only once. In this characterization experiment, 200,000 realizations of the 9 -bit input vector ( $\left.\boldsymbol{c}_{\mathrm{in}}, \boldsymbol{a}\right)$ are randomly generated by Matlab. All input bits are of uniform distribution and independent of each other. Every 5 ns , a newly generated realization of $\left(\boldsymbol{c}_{\mathrm{in}}, \boldsymbol{a}\right)$ is fed to the blocks, and at the same time, their outputs are sampled. Finally, the required conditional probabilities are estimated using (5) and (6).

When characterizing a probabilistic block, noise filtering effect of logic gates has to be taken into account. It was observed in [12] that an output signal of a logic gate is often less noisy than the corresponding (noisy) input signal. The main reason for this appears to be the so-called noise filtering effect discussed in Section VI.B of [13]. In short, noise filtering occurs when the duration of a noise pulse is shorter than the propagation delay of a gate. It was observed in [12] that it has prominent effect on the accuracy of characterization of a probabilistic circuit.


Fig. 8. Error-rates obtained by (2) and (3), and by HSPICE.

A practical way to deal with noise filtering effect is to use a buffer to approximate the output loads of an output terminal $[5,9,10,12]$. The buffer is connected to a block's output, and output samples are taken after this buffer. A buffer should have a suitable size, in order to accurately model the actual loads that will be connected to the block when it is chained with other blocks to form a PCSLA.

We also observe that different supply voltages require slightly different buffer sizes. This is because the supply voltage of a logic gate has an effect on its propagation delay, which in turn affects the strength of noise filtering. However, for simplicity, in this paper we only use a single buffer size for all of the voltages $0.8-1.2 \mathrm{~V}$.

The experiment results are shown in Fig. 8, which shows the predicted and simulated error-rates of the sumbits. "Decomposition 1" in Fig. 8 refers to the PCSLA of Fig. 3 (a). We see that the predictions are fairly close to the simulated error-rates for various supply voltages. For clarity, the same results are also reproduced in Table II. The numbers of errors made by the sum-bits in the simulation are shown in the Sim rows. For example, for 0.8 V , 3,677 out of 200,000 samples taken at the seventh sum-bit (S7) are incorrect. The predicted number is 3,671 , which is shown in the Pred-1 row. This number is given by multiplying the predicted error-rate by 200,000 and rounding to the nearest integer.

On average over all bit positions, a predicted error-rate is within 1-6\% of the corresponding simulated value when the supply voltage is within $0.8-1.1 \mathrm{~V}$. Whereas, for 1.2 V , the deviations are on average $11 \%$ of the simulated values. The relatively large deviation for 1.2 V is due to the fact that 1.2 V is the nominal voltage of the Synopsys 90 nm technology and hence errors rarely occur when the probabilities in (4) were estimated in the characterization process. The estimated probabilities are therefore somewhat less accurate.

## C. Prediction with Another Decomposition Approach

We now use another way to decompose a PCSLA. We decompose a PCSLA into three types of blocks shown in Fig. 5. Hence, besides Fig. 3 (a), the PCSLA can also be

TABLE II
Number of errors (out of 200,000 samples) predicted by the proposed equations and simulated by HSPICE.

| V | - | S 1 | S 2 | S 3 | S 4 | S 5 | S 6 | S7 | S8 | S9 | S10 | S11 | S12 | S13 | S14 | S15 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| S16 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.8 | Sim | 2150 | 2662 | 3226 | 3515 | 4311 | 3770 | 3677 | 3716 | 4543 | 3920 | 3709 | 3755 | 4550 | 3832 | 3622 |
| 3850 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.8 | Pred-1 | 2218 | 2762 | 3234 | 3668 | 3994 | 3643 | 3671 | 3882 | 4099 | 3696 | 3697 | 3895 | 4106 | 3699 | 3698 |
| 3896 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.8 | Pred-2 | 2233 | 2704 | 2871 | 2963 | 3719 | 3443 | 3239 | 3146 | 3809 | 3488 | 3262 | 3157 | 3815 | 3491 | 3263 |
| 3158 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.9 | Sim | 1108 | 1263 | 1387 | 1525 | 1895 | 1622 | 1529 | 1548 | 1890 | 1687 | 1587 | 1633 | 1925 | 1662 | 1498 |
| 1666 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.9 | Pred-1 | 1157 | 1275 | 1383 | 1547 | 1866 | 1629 | 1559 | 1635 | 1909 | 1651 | 1570 | 1640 | 1912 | 1652 | 1571 |
| 1640 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 0.9 | Pred-2 | 1117 | 1318 | 1366 | 1403 | 1809 | 1663 | 1539 | 1489 | 1852 | 1684 | 1549 | 1494 | 1855 | 1686 | 1550 |
| 1495 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1.0 | Sim | 503 | 570 | 599 | 652 | 749 | 698 | 625 | 624 | 777 | 714 | 651 | 673 | 808 | 679 | 614 |
| 1.0 | Pred-1 | 533 | 564 | 564 | 618 | 768 | 681 | 623 | 647 | 783 | 689 | 627 | 649 | 784 | 689 | 627 |
| 1.0 | Pred-2 | 520 | 591 | 625 | 647 | 805 | 734 | 697 | 683 | 823 | 743 | 701 | 685 | 824 | 744 | 701 |
| 1.1 | Sim | 207 | 235 | 263 | 261 | 295 | 270 | 270 | 247 | 292 | 286 | 237 | 272 | 323 | 261 | 233 |
| 1.1 | Pred-1 | 222 | 238 | 229 | 253 | 314 | 284 | 252 | 264 | 320 | 287 | 253 | 265 | 320 | 287 | 254 |
| 1.1 | Pred-2 | 197 | 230 | 245 | 256 | 313 | 288 | 274 | 271 | 321 | 292 | 276 | 272 | 321 | 292 | 276 |
| 1.2 | Sim | 85 | 86 | 89 | 105 | 100 | 84 | 116 | 100 | 114 | 109 | 90 | 102 | 123 | 107 | 94 |
| 1.2 | Pred-1 | 87 | 84 | 102 | 110 | 114 | 98 | 109 | 113 | 116 | 99 | 109 | 114 | 116 | 99 | 109 |
| 1.2 | Pred-2 | 81 | 95 | 101 | 105 | 137 | 123 | 115 | 112 | 141 | 125 | 116 | 113 | 141 | 125 | 116 |

represented as the cascade circuit illustrated in Fig. 3 (b). Because there are three different types of blocks, three rounds of characterization are required, each round for each individual block. We select the same parameters as those used previously, i.e., 200,000 input realizations, uniformly distributed and independent input bits, 5 ns sampling time, and 0.2 V noise RMS.

The predictions by (2) and (3) are shown in Fig. 8 (decomposition 2). The results are also shown in the Pred-2 row of Table II. The predicted error-rates are, again, fair close to those simulated by HSPICE. Nevertheless, they are not as accurate as those obtained by the first decomposition method presented previously (Pred-1). On average over all bit positions, a predicted error-rate is within 3$7 \%$ of the corresponding simulated value when the supply voltage is within $0.9-1.1 \mathrm{~V}$. Whereas, for 0.8 and 1.2 V , the deviations are on average $12 \%$ and $18 \%$, respectively, of the simulated values.

A possible explanation is that the influence of supply voltage on the noise filtering effect of logic gates now becomes more significant. Since we have used only one buffer size for all supply voltages, the predictions would be accurate for some voltages (e.g., $0.9-1.1 \mathrm{~V}$ ) but less accurate for other voltages. If our speculation is correct, then one can use different buffer sizes for different supply voltages in order to improve the accuracy.

## V. Conclusion

The method of [5] has been used to predict error-rates of a probabilistic carry-select adder (PCSLA). Experimental results suggested that the predictions were close to errorrates simulated by HSPICE simulations. These results provided evidence that the error-rate prediction method is versatile and applicable to some non-trivial circuits.

We further suggest that, by appropriately grouping circuit components into various blocks, the error-rate prediction method can handle not only cascade circuits, but also circuits with more complex structure. This has been demonstrated with a PCSLA, in which a multi-bit adder circuit with a seemingly non-cascade structure was trans-
formed into a cascade circuit, which was then handled successfully by the error-rate prediction method.

## References

[1] K. V. Palem, "Energy aware computing through probabilistic switching: A study of limits," IEEE Transactions on Computers, vol. 54, no. 9, pp. 1123-1137, 2005.
[2] J. George, B. Marr, B. E. S. Akgul, and K. V. Palem, "Probabilistic arithmetic and energy efficient embedded signal processing," Proceedings of 2006 CASES, pp. 158-168.
[3] M. S. K. Lau, K. V. Ling, and Y. C. Chu, "Energy-aware probabilistic multiplier: Design and analysis," Proceedings of 2009 CASES, pp. 281-290.
[4] L. N. Chakrapani, K. K. Muntimadugu, A. Lingamneni, J. George, and K. V. Palem, "Highly energy and performance efficient embedded computing through approximately correct arithmetic: A mathematical foundation and preliminary experimental validation," Proceedings of 2008 CASES, pp. 187-196.
[5] M. S. K. Lau, K. V. Ling, A. Bhanu, and V. J. Mooney III, "Error-rate prediction for a class of probabilistic circuits with applications to carry-skip adders," submitted to the 16th ASPDAC (obtainable from http://web.me.com/mark.lau).
[6] S. Krishnaswamy, G. F. Viamontes, I. L. Markov, and J. P. Hayes, "Probabilistic transfer matrices in symbolic reliability analysis of logic circuits," ACM Transactions on Design Automation of Electronic Systems, vol. 13, no. 1, Article 8, 2008.
[7] A. Chandrakasan, W. J. Bowhill, and F. Fox, Eds., Design for High-Performance Microprocessor Circuits. IEEE Press, 2000.
[8] J. M. Rabaey, A. Chandrakasan, and B. Nikolić, Digital Integrated Circuits: A Design Perspective, 3rd ed. Prentice Hall, 2003.
[9] M. S. K. Lau, K. V. Ling, Y. C. Chu, and A. Bhanu, "Modeling of probabilistic ripple-carry adders," Proceedings of 2010 DELTA, pp. 201-206.
[10] ——, "A general mathematical model of probabilistic ripplecarry adders," in Proceedings of 2010 DATE, CD-ROM.
[11] S. Ross, A First Course in Probability, 7th ed. Prentice Hall, 2006.
[12] A. Bhanu, M. S. K. Lau, K. V. Ling, V. J. Mooney III, and A. Singh, "A more precise model of noise based CMOS errors," Proceedings of 2010 DELTA, pp. 99-102.
[13] P. Korkmaz, B. E. S. Akgul, and K. V. Palem, "Energy, performance, and probability tradeoffs for energy-efficient probabilistic CMOS circuits," IEEE Transactions on Circuits and Systems I, vol. 55, no. 8, pp. 2249-2262, 2008.


[^0]:    *Also with Georgia Institute of Technology.

